Population-based Clinical Practice Research Datalink study using algorithm modelling to identify the true burden of hidradenitis suppurativa*

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Summary

Background Epidemiology data regarding hidradenitis suppurativa (HS) are conflicting and prevalence estimates vary 80-fold, from 0.05% in a population-based study to 4%.

Objectives To assess the hypothesis that previous population-based studies underestimated true HS prevalence by missing undiagnosed cases.

Methods We performed a population-based observational and case–control study using the U.K. Clinical Practice Research Datalink (CPRD) linked to hospital episode statistics data. Physician-diagnosed cases in the CPRD were identified from specific Read codes. Algorithms identified unrecognized ‘proxy’ cases, with at least five Read code records for boils in flexural skin sites. Validation of proxy cases was undertaken with general practitioner (GP) questionnaires to confirm criteria-diagnosed cases. A case–control study assessed disease associations.

Results On 30 June 2013, 23,353 physician-diagnosed HS cases were documented in 4,364,308 research-standard records. In total, 68,890 proxy cases were identified, reduced to 10,146 criteria-diagnosed cases after validation, extrapolated from 107 completed questionnaires (61% return rate). Overall point prevalence was 0.77% [95% confidence interval (CI) 0.76–0.78%]. An additional 18,417 cases had a history of one to four flexural skin boils. In physician-diagnosed cases, odds ratios (ORs) for current smoker and obesity (body mass index > 30 kg m⁻²) were 3.61 (95% CI 3.44–3.79) and 3.29 (95% CI 3.14–3.45). HS was associated with type 2 diabetes, Crohn disease, hyperlipidaemia, acne and depression, and not associated with ulcerative colitis or polycystic ovary syndrome.

Conclusions Contrary to results of previous population-based studies, HS is relatively common, with a U.K. prevalence of 0.77%, one-third being unrecognized, criteria-diagnosed cases using the most stringent disease definition. If individuals with probable cases are included, HS prevalence rises to 1.19%.

What’s already known about this topic?

• Previous population-based studies suggest that the prevalence of hidradenitis suppurativa (HS) may be as low as 0.05% but existing studies have not quantified undiagnosed cases.
Hidradenitis suppurativa (HS), also known as acne inversa, is an inflammatory skin disease producing functional impairment similar in magnitude to cardiovascular disease, type 2 diabetes and renal failure, because of pain, purulent discharge and scarring.1,2 HS is defined clinically by characteristic lesions occurring chronically in flexural locations.3,4 Consensus disease definitions differ slightly in the chronicity element, ranging from a history of two to five boils in the axillae or groins.5,6

The prevalence of HS remains controversial, with estimates ranging from 0·05% from an analysis of patient insurance claims in the U.S.A.7 to 4% when young adult women were examined in person.8 An accurate prevalence figure is required to quantify the unmet need regarding care of people with HS to assess disease burden and plan healthcare provision. In particular, radical surgery or biological therapies are needed for severe disease that are costly and often require multidisciplinary teams; for example, in the U.K., the National Institute for Health and Care Excellence recently approved adalimumab for moderate-to-severe HS.9

One explanation for the 80-fold difference in prevalence estimates currently available is lack of recognition of HS, in the context that the average diagnostic delay is 7 years.10 The aim of our study was to provide the most accurate estimate to date of HS prevalence by quantifying physician-diagnosed cases, and also unrecognized HS cases meeting the most stringent consensus disease definition for HS. A primary care data source was chosen because the majority of HS care takes place within a primary care setting. Our second aim was to assess HS disease associations with the largest case–control study performed to date. Accurate epidemiological data is critical at a time when treatment for HS is dramatically changing and now includes biological therapies.

Patients and methods

Our manuscript was prepared in accordance with RECORD (Reporting of Studies Conducted using Observational Routinely-collected Data)11 (for the RECORD checklist see Appendix S1 in the Supporting Information). The Clinical Practice Research Datalink (CPRD) contains routinely recorded, pseudonymized data from participating primary care practices throughout the U.K.12 Most diagnoses made in secondary care are transcribed into the CPRD by primary care staff. As of January 2015, it contained records from more than 15·6 million patients registered at 684 practices. Patient data are labelled by the CPRD as being of acceptable research quality if the patient has been permanently registered at their general practice and has internally consistent records with regard to age, sex, registration and event dates. Approximately 75% of contributing practices based in England (representing 60% of the entire dataset) participate in a CPRD linkage scheme by which patient-level data are linked to the hospital episode statistics (HES) for England. The National Health Service (NHS) number, date of birth, sex and postal code of nondissenting patients are conveyed from the practice to a trusted third party and matched with equivalent identifiers in the target dataset. Only those linkages involving at least two exact matches, which must include an exact match on NHS number, are retained.

The study population was selected from patients with data of acceptable research quality. Physician-diagnosed cases of HS were captured by two specific diagnostic Read codes, M25y100 (Hidradenitis) and M25y111 (Hidradenitis suppurativa), or, in HS data, by the International Classification of Disease (ICD)-10 code L73.2 (Hidradenitis suppurativa). To capture undiagnosed cases, CPRD Read code algorithms were created to identify patients attending primary care for multiple skin boil consultations (Table 1). An algorithm hierarchy was used with different degrees of stringency for each subalgorithm, based on the HS disease definition (Fig. 1). Some primary care practitioners may include HS patients as being of acceptable research quality if the patient has been permanently registered at their general practice and has internally consistent records with regard to age, sex, registration and event dates. Approximately 75% of contributing practices based in England (representing 60% of the entire dataset) participate in a CPRD linkage scheme by which patient-level data are linked to the hospital episode statistics (HES) for England. The National Health Service (NHS) number, date of birth, sex and postal code of nondissenting patients are conveyed from the practice to a trusted third party and matched with equivalent identifiers in the target dataset. Only those linkages involving at least two exact matches, which must include an exact match on NHS number, are retained.

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not record every skin boil consultation, particularly if the patient also attends with other health issues. To capture these potential cases, treatment codes were used to identify patients receiving multiple short courses of skin-directed antibiotic therapy, such as flucloxacillin, in combination with skin boil Read codes and in the absence of any other indication for the antibiotic, such as eczema, cellulitis or skin ulcers. Prior to validation, the cases of these patients were referred to as ‘proxy cases’. HES data were not used to identify proxy cases to avoid double-counting boils recorded in both data sources.

Validation

Validation of proxy cases was carried out by sending questionnaires to the primary care physicians of a subset of patients identified from the algorithms. The questionnaire (Table 2) requested confirmation of characteristic skin lesions in flexural locations on at least five separate occasions, in keeping with the most stringent consensus definition of the disease.6

A total of 176 questionnaires were sent, with the overall number being determined by our study budget in the context that CPRD charges a fee for each questionnaire issued. The number of questionnaires generated to validate each subalgorithm was determined by the number of additional proxy cases identified through each, with a minimum of 10 questionnaires being issued for subalgorithms that captured relatively few cases.

Analysis

Point prevalence on 30 June 2013 was calculated using physician-diagnosed cases and criteria-diagnosed cases as the numerator and the 2013 mid-year population of live patients with research-standard records in CPRD as the denominator. The number of criteria-diagnosed cases was calculated by multiplying the number of proxy cases identified by each subalgorithm with the proportion of cases validated from the relevant set of GP questionnaires.

Annual incidence rates for physician-diagnosed cases were calculated from 1988, when up-to-research-standard practices first contributed to CPRD, until 2013. We used the Brameld–Holman backcasting method, which implements an actuarial retrograde survival curve to derive correction factors for the over-ascertainment of first events for individuals with only a limited duration of CPRD data collection prior to the event.13

The Brameld–Holman method accounts for the prevalent pool effect without the need for a clearance period and avoids data being discarded.

Case–control

Cases and controls were matched in a 1 : 1 ratio based on age, sex and registration at the same primary care practice. The associations between HS and environmental factors and other medical conditions were estimated using conditional logistic regression techniques. Two separate association analyses were conducted for physician-diagnosed cases compared with matched controls and for proxy cases (including criteria-diagnosed HS cases and those proxy cases who did not meet the validation criteria) compared with a further set of matched controls. Univariable associations were reported, and also multivariable analyses that examined combinations of variables demonstrating association in the univariable analysis with a
Results

Point prevalence

On 30 June 2013 there were 21 575 patients with physician-diagnosed HS in the CPRD dataset and a further 1778 from the HES dataset; a total of 23 353 individuals with physician-diagnosed HS. Our Read and treatment code algorithms captured a further 68 890 proxy cases in the CPRD; the breakdown for each subalgorithm is shown in Table 3. Of the 176 GP validation questionnaires sent, 107 (61%) were fully completed and returned; completion rates for each subalgorithm are in Table 3.

The proportions of questionnaires with 'yes' responses to the first three questions, satisfying the most stringent disease definition of at least five flexural skin boils, are documented in Table 3; from these, 10 146 criteria-diagnosed cases were extrapolated. Table 3 also documents the proportions within each subalgorithm confirmed to have had one to four flexural boils with 'yes' responses to questions one and two only, representing a group of probable HS cases. Overall, only one validation questionnaire confirmed the presence of scarring with a 'yes' answer to question four.

Using the denominator of 4 364 308 research-standard patients in the CPRD, the prevalence of physician-diagnosed and criteria-diagnosed HS cases on 30 June 2013 was 7.7 per 1000 [95% confidence interval (CI) 7.6–7.8]. When the 18 417 patients with one to four flexural skin boils were included, the upper limit of probable HS prevalence was 11.9 per 1000 (95% CI 11.8–12.0).

The demographic profile of patients with physician-diagnosed HS is shown in Figure 2. Peak prevalence of 15.1 per 1000 occurred in the fifth decade of life and the mean female : male ratio across all the age groups was 2.9 : 1.

Incidence of physician-diagnosed cases

The Brameld–Holman actuarial retrograde survival curve used for the backcasting correction for incidence rates is in Figure S1 (see Supporting Information), with the resulting correction factors reducing the number of new physician-diagnosed cases prior to 1994. The mean annual incidence rate for physician-diagnosed cases from 1996 to 2013 was 28.3 per 100 000 person-years (Table S1; see Supporting Information). Prior to 1995, despite the backcasting correction, annual incidence rates became progressively higher, ranging from 35.7 per 100 000 person-years in 1995 to 122.1 in 1988.

Case–control

Descriptive statistics for the occurrence of environmental factors and other medical conditions in cases and controls are in Table 4. Obesity [body mass index (BMI) > 30 kg m⁻²], current smoking and type 2 diabetes mellitus were strongly linked with physician-diagnosed cases of HS, with odds ratios (ORs) of 3.29 (95% CI 3.14–3.45), 3.61 (95% CI 3.44–3.79) and 3.39 (3.09–3.71), respectively, in the univariable analysis (Table 4). There was a significant association between HS and Crohn disease but not ulcerative colitis. The link with Crohn disease was not stronger for smokers than nonsmokers (P = 0.26 for the interaction test between Crohn disease and never or past/current smoker). Associations were also found

Table 3 Subalgorithm validation from completed questionnaires returned by primary care practitioners

<table>
<thead>
<tr>
<th>Subalgorithm</th>
<th>Questionnaires sent</th>
<th>Questionnaires completed</th>
<th>Proxy cases prior to validation</th>
<th>Proportion validated with ≥ 5 flexural boils</th>
<th>Extrapolated number of criteria-diagnosed casesa</th>
<th>Proportion validated with 1–4 flexural boils</th>
<th>Extrapolated additional probable casesb</th>
</tr>
</thead>
<tbody>
<tr>
<td>1a</td>
<td>10</td>
<td>6</td>
<td>650</td>
<td>0.67</td>
<td>435</td>
<td>0</td>
<td>0</td>
</tr>
<tr>
<td>1b</td>
<td>10</td>
<td>6</td>
<td>1908</td>
<td>0.33</td>
<td>629</td>
<td>0.17</td>
<td>325</td>
</tr>
<tr>
<td>1c</td>
<td>10</td>
<td>7</td>
<td>3120</td>
<td>0.57</td>
<td>1778</td>
<td>0.29</td>
<td>905</td>
</tr>
<tr>
<td>2a</td>
<td>10</td>
<td>3</td>
<td>2083</td>
<td>0.33</td>
<td>687</td>
<td>0.34</td>
<td>708</td>
</tr>
<tr>
<td>2b</td>
<td>59</td>
<td>38</td>
<td>27 492</td>
<td>0.13</td>
<td>3573</td>
<td>0.29</td>
<td>7973</td>
</tr>
<tr>
<td>3</td>
<td>0</td>
<td>0</td>
<td>0</td>
<td>0</td>
<td>0</td>
<td>0</td>
<td>0</td>
</tr>
<tr>
<td>4</td>
<td>10</td>
<td>5</td>
<td>594</td>
<td>0</td>
<td>0</td>
<td>0.20</td>
<td>118</td>
</tr>
<tr>
<td>5a</td>
<td>14</td>
<td>10</td>
<td>7050</td>
<td>0.10</td>
<td>705</td>
<td>0.60</td>
<td>4229</td>
</tr>
<tr>
<td>5b</td>
<td>53</td>
<td>32</td>
<td>25 993</td>
<td>0.09</td>
<td>2339</td>
<td>0.16</td>
<td>4159</td>
</tr>
<tr>
<td>Total</td>
<td>176</td>
<td>107</td>
<td>68 890</td>
<td>0.14</td>
<td>10 146</td>
<td>18 417</td>
<td></td>
</tr>
</tbody>
</table>

aCriteria-diagnosed cases are defined as those individuals confirmed by their primary care physician to have a history of at least five flexural skin boils; badditional probable cases are defined as those individuals confirmed by their primary care physician to have a history of between one and four flexural skin boils.
between HS and hyperlipidaemia, pilonidal sinus, acne vulgaris, depression and hypertension, but not with inflammatory arthritis, polycystic ovary syndrome, psoriasis or Alzheimer disease (Table 4). Multivariable analysis confirmed the univariable analysis results in all cases (Table S2; see Supporting Information).

**Discussion**

Our HS prevalence figure of 0.77% is derived from a population of 4.3 million U.K. residents with research-standard medical records. Importantly, our study has identified patients with previously undiagnosed HS using the most stringent diagnostic criteria available. We have also identified an upper limit of probable HS prevalence of 1.19%, if those with a history of one to four flexural skin boils on separate occasions are included. This is in the context that patients may not see their GP for every new skin boil, GPs may not record each skin boil, particularly if the consultation also covers other medical issues, and HS disease definitions vary in the number of flexural skin boils required, from two to five.5,6

Our two prevalence figures span the 0.97% estimate for European prevalence derived from a self-reported questionnaire study of 10,000 members of the French population that

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**Table 4** Case–control study univariable analysis

<table>
<thead>
<tr>
<th></th>
<th>Controls, n (%)</th>
<th>Proxy cases, n (%)</th>
<th>Physician-diagnosed cases, n (%)</th>
<th>P-value</th>
<th>OR (95% CI)</th>
<th>P-value</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Obesity</strong> (body mass index &gt; 30 kg m⁻²)</td>
<td>19 862 (21.2)</td>
<td>29 916 (42.8)</td>
<td>11 145 (46.4)</td>
<td>&lt; 0.001</td>
<td>2.77 (2.70–2.85)</td>
<td>&lt; 0.001</td>
</tr>
<tr>
<td><strong>Type 2 diabetes mellitus</strong></td>
<td>4321 (4.6)</td>
<td>9530 (13.7)</td>
<td>2158 (9.0)</td>
<td>&lt; 0.001</td>
<td>3.32 (3.18–3.47)</td>
<td>&lt; 0.001</td>
</tr>
<tr>
<td><strong>Smoking status</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Never</td>
<td>48 431 (51.6)</td>
<td>28 751 (41.2)</td>
<td>7828 (32.6)</td>
<td>1</td>
<td>1</td>
<td></td>
</tr>
<tr>
<td>Past</td>
<td>12 671 (13.5)</td>
<td>10 523 (15.1)</td>
<td>3260 (13.6)</td>
<td>1.44 (1.39–1.49)</td>
<td>&lt; 0.001</td>
<td>1.93 (1.81–2.05)</td>
</tr>
<tr>
<td>Current</td>
<td>22 712 (24.2)</td>
<td>27 748 (39.7)</td>
<td>12 347 (51.4)</td>
<td>2.18 (2.13–2.24)</td>
<td>&lt; 0.001</td>
<td>3.61 (3.44–3.79)</td>
</tr>
<tr>
<td><strong>Crohn disease</strong></td>
<td>187 (0.2)</td>
<td>310 (0.4)</td>
<td>122 (0.5)</td>
<td>2.21 (1.81–2.69)</td>
<td>&lt; 0.001</td>
<td>2.65 (1.89–3.72)</td>
</tr>
<tr>
<td><strong>Ulcerative colitis</strong></td>
<td>129 (0.1)</td>
<td>78 (0.1)</td>
<td>42 (0.2)</td>
<td>0.82 (0.61–1.11)</td>
<td>&lt; 0.001</td>
<td>0.2</td>
</tr>
<tr>
<td><strong>Pilonidal sinus</strong></td>
<td>293 (0.3)</td>
<td>887 (1.3)</td>
<td>444 (1.9)</td>
<td>4.21 (3.62–4.89)</td>
<td>&lt; 0.001</td>
<td>5.61 (4.41–7.12)</td>
</tr>
<tr>
<td><strong>Acne vulgaris</strong></td>
<td>5828 (6.2)</td>
<td>7858 (11.3)</td>
<td>2885 (12.0)</td>
<td>2.17 (2.09–2.27)</td>
<td>&lt; 0.001</td>
<td>1.77 (1.66–1.88)</td>
</tr>
<tr>
<td><strong>Depression</strong></td>
<td>16 009 (17.1)</td>
<td>16 968 (24.3)</td>
<td>6318 (26.3)</td>
<td>1.65 (1.60–1.69)</td>
<td>&lt; 0.001</td>
<td>1.69 (1.62–1.77)</td>
</tr>
<tr>
<td><strong>Hyperlipidaemia</strong></td>
<td>2664 (2.8)</td>
<td>3344 (4.8)</td>
<td>791 (3.3)</td>
<td>1.59 (1.50–1.69)</td>
<td>&lt; 0.001</td>
<td>1.79 (1.59–2.02)</td>
</tr>
<tr>
<td><strong>Hypertension</strong></td>
<td>8090 (8.6)</td>
<td>8596 (12.3)</td>
<td>1882 (7.8)</td>
<td>1.40 (1.35–1.45)</td>
<td>&lt; 0.001</td>
<td>1.42 (1.32–1.53)</td>
</tr>
<tr>
<td><strong>Inflammatory arthritis</strong></td>
<td>790 (0.8)</td>
<td>634 (0.9)</td>
<td>201 (0.8)</td>
<td>1.06 (0.94–1.18)</td>
<td>0.35</td>
<td>1.06 (0.87–1.00)</td>
</tr>
<tr>
<td><strong>Polycystic ovary syndrome</strong></td>
<td>377 (0.4)</td>
<td>252 (0.4)</td>
<td>181 (0.8)</td>
<td>1.12 (0.94–1.34)</td>
<td>0.22</td>
<td>1.19 (0.96–1.49)</td>
</tr>
<tr>
<td><strong>Psoriasis</strong></td>
<td>879 (0.9)</td>
<td>671 (1.0)</td>
<td>247 (1.0)</td>
<td>1.03 (0.92–1.13)</td>
<td>0.62</td>
<td>1.09 (0.91–1.31)</td>
</tr>
<tr>
<td><strong>Alzheimer disease</strong></td>
<td>102 (0.1)</td>
<td>87 (0.1)</td>
<td>16 (0.1)</td>
<td>0.95 (0.71–1.27)</td>
<td>0.71</td>
<td>1.60 (0.73–3.53)</td>
</tr>
</tbody>
</table>

*Based on 93 869 control individuals, 69 842 proxy cases including criteria-diagnosed hidradenitis suppurativa (HS) cases and those who did not meet the validation criteria, and 24 027 physician-diagnosed cases of HS. Some non-prevalent cases, meeting diagnostic criteria after 30 June 2013, are included. Smoking status was missing from a total of 10 055 records, including 2820 proxy cases and 592 physician-diagnosed cases and missing data is included in the denominator for calculation of percentages. Body mass index data were missing from 12% of all records and, again, missing data are included in the denominator for calculation of percentages.*
is often cited as the most accurate current estimate. In the French study, cases were identified by positive responses to one question within a general skin disease survey and the morphology of skin lesions was not assessed by healthcare professionals.

One of the largest population-based HS investigations performed prior to our study used data from the Rochester Epidemiology Project (REP), which maintains a centralized electronic medical record database for the 144,000 people living in Olmsted County Minnesota, U.S.A. Identification of HS cases used codes including synonyms of 'hidradenitis', and also the terms 'infection, sweat gland or inflammation, sweat gland'. Validation by chart review identified 178 HS cases, giving a prevalence of 0.13%. More recently, interrogation of pooled insurance and self-pay databases covering 48 million people in the U.S.A. using ICD-9 coding for HS gave a point prevalence of 0.10%. In both studies, identification of undiagnosed cases by considering multiple flexural skin boil presentations was not attempted. Part of the explanation for the lower U.S.A. prevalence figures than in our data is failure to capture undiagnosed cases. In addition, recognition and coding of HS by healthcare professionals will vary from region to region and variations in health systems may influence whether people with HS persevere with seeking medical attention for their condition, an issue requiring further investigation.

Our finding of peak HS prevalence in women in their fourth and fifth decades of life is in agreement with the French self-reported questionnaire study, in which the female : male ratio was 2.7 : 1 and the mean age of patients was 43.2 years. Our peak prevalence finding of 1.51% in women in their fifth decade is lower than the 4% prevalence found by examination of a group of female healthcare professionals and those attending for benign skin lesions in Denmark, but the study had a relatively low sample size of 100, in which there were four cases.

Our mean annual incidence of physician-diagnosed HS cases of 28.3 per 100,000 person-years from 1996 to 2013 is more than four times higher than the 6.0 per 100,000 person-years figure found using REP data for Olmsted County, Minnesota. The difference is in keeping with the higher prevalence figure for physician-diagnosed cases found in our study. The more recent study of the U.S.A. population using pooled insurance and self-pay databases found a HS incidence of 11.4 (95% CI 11.1–11.8) per 100,000 population, double the incidence of the REP-derived estimate. Our calculated incidence rates were much higher in the first few years of available CPRD data, despite the backcasting correction to negate any prevalent pool effect. The observation is likely explained by bulk computerization of patient paper records during this period, producing an artefact in this result.

Our case–control study is entirely consistent with other, smaller studies previously performed. A recent systematic review of nine relevant studies containing 6174 patients and 24,993 controls confirmed an association between HS and obesity (OR 3.45, 95% CI 2.20–5.38), in addition to an association with diabetes (OR 2.85, 95% CI 1.34–6.08), current smoking (OR 4.34, 95% CI 2.48–7.60) and past smoking (OR 6.34, 95% CI 2.41–16.68). Our finding that HS is associated with Crohn disease but not ulcerative colitis is in agreement with other studies, including a cross-sectional investigation of 3267 patients with HS and, in addition, we showed that the association between HS and Crohn disease is not modified by smoking status. The higher rates of smoking, type 2 diabetes mellitus, hypertension and hyperlipidaemia confirmed by our study are in keeping with the nearly doubled risk of cardiovascular-associated death in Danish patients with HS compared with controls (adjusted incidence rate ratio 1.95, 95% CI 1.42–2.67).

The strongest association demonstrated by our case–control study was between HS and pilonidal sinus, with an odds ratio of 5.61. The result is expected because both conditions are members of the 'follicular occlusion tetrad' and there is ongoing debate about whether pilonidal sinus may be a phenotypic variant of HS. We confirmed a link between HS and depression, with an odds ratio of 1.55, which is in keeping with a chronic, painful and socially isolating disease and in agreement with a study of psychiatric comorbidities in 3207 patients with HS (adjusted OR for depression 1.7, 95% CI 1.3–2.1). However, we did not assess the timing of disease associations in our study and so cannot comment further on possible causation.

Debate continues regarding a possible link between HS and inflammatory arthritis. A syndrome known as PAPASH (pyoderma gangrenosum, acne, psoriasis, arthritis and supplicative hidradenitis) has been proposed but we did not find evidence of an association between HS and inflammatory arthritis or psoriasis. Our finding of no link between HS and polycystic ovary syndrome disagrees with a case–control study based on 2292 patients in the U.S.A. with at least one billing code for HS, which found an odds ratio of 13.7 (95% CI 4.00–47.3). However, the study found a HS prevalence of only 0.08%, reflecting limited case-capture from billing codes, and so case ascertainment bias may have affected the results. We specifically checked for any link between HS and Alzheimer disease because of reports of loss of function gamma secretase gene mutations in some Han Chinese patients with HS and a small proportion of European families with HS. Missense gamma secretase gene mutations are implicated in familial Alzheimer disease. We found no association between HS and Alzheimer disease, in agreement with a recent case–control study from Denmark, although it is possible that our relatively young cohort of patients might subsequently develop dementia later in life.

Considering our study limitations, some relevant data is missing within the CPRD and it is likely that these will not be missing at random because patients who are cases are likely to have greater completeness because of their condition, increasing GP contacts and providing more opportunity for the data items to be recorded. Coding imperfections may affect results, mitigated by our validation of undiagnosed cases using the GP questionnaire. Study budget limitations, rather than statistical...
considerations, restricted the number of validation questionnaires that were sent and so a sampling error is possible.

We did not validate physician-diagnosed HS cases because false positives are unlikely in the context that HS is a clinically diagnosed, characteristic skin disease with a low chance of being confused with other skin conditions. Review of medical records for patients who had received at least two ICD-9 codes for ‘hidradenitis’ in the Massachusetts General Hospital database was previously conducted, in the context that the 705-83 code includes neutrophilic eccrine hidradenitis, and recurrent palmoplantar hidradenitis, in addition to HS.29 Confirmation of HS was obtained in 89-6% of cases, while inclusion of the terms ‘hidradenitis’ or ‘HS’ in the medical record gave positive predictive values of 99-6% (95% CI 98.9–99.9%) and 100% (95% CI 95.8–100%), respectively, for identification of chart-verified HS. To minimize any risk of false positives we used only the Read codes for ‘hidradenitis’ and ‘hidradenitis suppurativa’ in our search for physician-diagnosed cases in the CPRD and used the specific ICD-10 code for ‘hidradenitis suppurativa’, L73-2, when searching the HES data.

Validation in person would have been ideal but is not possible within the ethical parameters of the CPRD. Validation using a GP questionnaire to check for the consensus definition of five flexural skin boils would be prone to false negatives. This is because five skin boil consultations may not be recorded as separate entries in the medical record in those with a known diagnosis. For the same reason, we did not analyse the performance of our algorithms in detecting physician-diagnosed cases. Descriptive statistics comparing controls with proxy cases and diagnosed cases support a conclusion that the proxy group contains mainly patients with HS and with proxy cases and diagnosed cases support a conclusion that HS is a common condition, highlighting the relative lack of HS clinical trial evidence to date. For example, the recent HS Cochrane review included only 12 randomized controlled trials, involving only 615 participants, one-seventh of the number of trials and participants included in the updated Cochrane review of vitiligo, a condition with a similar prevalence.30,31 Availability of an accurate prevalence figure is one of the key foundations to drive further research in the field of HS.

References


Supporting Information

Additional Supporting Information may be found in the online version of this article at the publisher’s website:

Appendix S1 Completed RECORD (Reporting of Studies Conducted using Observational Routinely-collected Data) checklist.

Appendix S2 Study protocol (15_020R) approved by the Clinical Practice Research Datalink Independent Scientific Advisory Committee.

Fig S1. Brameld–Holman backcasting actuarial retrograde survival curve.

Table S1 Incidence of physician-diagnosed hidradenitis suppurativa cases in the U.K. from 1988 to 2013 including a backcasting correction for the prevalent pool effect.

Table S2 Case–control study multivariable analysis.

Powerpoint S1 Journal Club Slide Set.