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“What time is my next meal?” Delay-discounting individuals choose smaller portions under conditions of uncertainty

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Abstract

‘Dietary’ delay discounting is typically framed as a trade-off between immediate rewards and long-term health concerns. Our contention is that prospective thinking also occurs over shorter periods, and is engaged to select portion sizes based on the interval between meals (inter-meal interval; IMI). We sought to assess the extent to which the length of an IMI influences portion-size selection. We predicted that delay discounters would show ‘IMI insensitivity’ (relative lack of concern about hunger or fullness between meals). In particular, we were interested in participants’ sensitivity to an uncertain IMI. We hypothesized that when meal times were uncertain, delay discounters would be less responsive and select smaller portion sizes. Participants ($N=90$) selected portion sizes for lunch. In different trials, they were told to expect dinner at 5pm, 9pm, and either 5pm or 9pm (uncertain IMI). Individual differences in future-orientation were measured using a monetary delay-discounting task. Participants chose larger portions when the IMI was longer ($p < .001$). When the IMI was uncertain, delay-discounting participants chose smaller portions than the average portion chosen in the certain IMIs ($p < .05$). Furthermore, monetary discounting mediated a relationship between BMI and smaller portion selection in uncertainty ($p < .05$). This is the first study to report an association between delay discounting and IMI insensitivity. We reason that delay discounters selected smaller portions because they were less sensitive to the uncertain IMI, and overlooked concerns about potential future hunger. These findings are important because they illustrate that differences in discounting are expressed in short-term portion-size decisions and suggest that IMI insensitivity increases when meal timings are uncertain. Further research is needed to confirm whether these findings generalise to other populations.

Keywords: Chaotic eating, Impulsivity, Delay discounting, Meal planning, Portion size

39 **Introduction**

40 Impulsivity is a multidimensional construct that can be measured in various ways (Evdenden,
41 1999; Whiteside & Lynam, 2016). Delay discounting is a facet of impulsivity, referring to the
42 tendency to respond to the immediate rather than the long-term consequences of a decision
43 (Moeller, Barratt, Dougherty, Schmitz, & Swann, 2001). It is considered a behavioural-
44 economic index of impulsive decision-making (MacKillop et al., 2011). A non-future
45 oriented individual who discounts delayed rewards is often described as a ‘steep’ delay
46 discounter. Steep temporal discounting has been related to an unhealthy diet, overeating, and
47 obesity (Barlow, Reeves, McKee, Galea, & Stuckler, 2016; Kulendran et al., 2014;
48 Manwaring, Green, Myerson, Strube, & Wilfley, 2011; Rollins, Dearing, & Epstein, 2010).
49 Nevertheless, associations are often weak and unreliable (Appelhans et al., 2011; Eisenstein
50 et al., 2015; Hendrickson, Rasmussen, & Lawyer, 2015; Leitch, Morgan, & Yeomans, 2013;
51 Rasmussen, Lawyer, & Reilly, 2010; Stoeckel, 2013; Stojek, Fischer, Murphy, & MacKillop,
52 2014; Weller, Cook, Avsar, & Cox, 2008).

53 One explanation for these inconsistencies is that delay discounting can have multiple
54 effects on food decisions. By contrast, the role of temporal discounting is often framed
55 around a single proposition; that impulsive people overeat because they discount long-term
56 health consequences (Zhang & Rashad, 2008). For example, associations between
57 discounting and overconsumption are often attributed to a lack of concern for future weight
58 gain (Barlow, et al., 2016). This perspective stands at odds with research in both humans
59 (Gregorios-Pippas, Tobler, & Schultz, 2009; McClure, Ericson, Laibson, Loewenstein, &
60 Cohen, 2008; Tanaka et al., 2004) and non-human animals (Mazur, 2001; Shelley, 1993),
61 which shows that temporal discounting operates over much shorter delays of seconds and
62 minutes. Recent studies have found that humans also discount the value of food and drink at
63 intervals as short as thirty seconds (Hendrickson & Rasmussen, 2013; Lumley, 2016;

64 Rasmussen, et al., 2010). This indicates that people also discount short-term consequences of
65 dietary decisions, rather than just long-term concerns about health or weight gain. In the
66 present study we considered the prospect that dietary discounting occurs over an intermediate
67 time frame (hours rather than years) and is evident in the selection of portion sizes from one
68 meal to the next.

69 The majority of meals are planned in advance – people tend to select a portion to eat
70 and then clean their plate (Fay et al., 2011; Wilkinson et al., 2012). Portion size is often
71 governed by the ‘expected satiety’ of a food – a concern to select an amount that is sufficient
72 to stave off hunger (the desire to eat) in the interval between meals (Brunstrom & Rogers,
73 2009; Brunstrom, Shakeshaft, & Scott-Samuel, 2008). Anticipated meals timings probably
74 influence these decisions. However, no studies have systematically explored this
75 phenomenon and it remains unclear how monetary delay discounting relates to meal planning
76 in this context. To address these questions we explored the extent to which the length of an
77 inter-meal interval (IMI) influences lunchtime portion-size selection.

78 One possibility is that meal planning might be less evident in steeper discounters.
79 People plan their behaviours by evaluating the future consequences of a decision (da Matta,
80 Gonçalves, & Bizarro, 2012). However, impulsive decision-makers may fail to consider all
81 relevant information before making a choices (Verplanken & Sato, 2011). Given this logic,
82 we anticipated that steep delay discounters would be less concerned with the relative
83 consequences of a long or short IMI when making in-the-moment portion-size judgements.
84 Therefore, we reasoned that steep discounters would show ‘IMI insensitivity’, (a relative lack
85 of concern for potential hunger or fullness during the IMI) and have a smaller difference
86 between portion sizes chosen at a short and long IMI.

87 In addition, we are interested in the effects of an uncertain IMI. Traditionally, a
88 Westernised meal pattern comprises three primary meals; breakfast, lunch, and dinner.

89 However, sometimes the IMI is uncertain. Recently, there has been an increase in ‘chaotic
90 eating’ - snacking and eating meals at different times on different days (Samuelson, 2000;
91 Warde & Yates, 2016). Irregular eating is associated with having a higher BMI (Sierra-
92 Johnson et al., 2008) and is thought to be a contributing factor to high-energy intake and
93 weight gain (Berg & Forslund, 2015; Murata, 2000). Unsurprisingly, various dimensions of
94 impulsivity have been associated with chaotic eating behaviours, including opportunistic
95 snacking and a preference for snack foods (Fay, White, Finlayson, & King, 2015;
96 Nederkoorn, Houben, Hofmann, Roefs, & Jansen, 2010).

97 One possibility is that irregular meal times encourage impulsive behaviours because
98 they generate uncertainty. Uncertainty has been shown to increase delay discounting;
99 individuals discount future rewards more steeply when the delayed event is perceived to be
100 more risky or less certain (Baumann & Odum, 2012; Green & Myerson, 2010; Patak &
101 Reynolds, 2007). It is important to mention that these studies manipulated the likelihood of
102 an event occurring, rather than uncertainty around the exact timing of an event. We propose
103 that uncertainty about the timing of an event may also increase discounting. When IMIs are
104 certain, individuals can make predictions about future hunger or satiety. However, when
105 event timings are variable, it is harder to plan for the future (Greville & Buehner, 2010). On
106 this basis, uncertainty may increase discounting of information about future meal timings. To
107 protect against the potential for hunger, individuals who are sensitive to the future might
108 select larger portions when the IMI is uncertain. Conversely, steep discounters may be less
109 responsive. Hence, we hypothesized that when meal timings were uncertain, steep delay
110 discounters would select portion sizes that are smaller the average of those chosen when meal
111 times were certain. We considered evidence for this hypothesis by systematically
112 manipulating the certainty of an IMI.

113 In the present study we measured portion selection in response to information about
114 the IMI. Participants chose lunch portions in three different conditions; two where the IMI
115 was ‘certain’ (dinnertime at 5pm and 9pm), and one where the IMI was ‘uncertain’
116 (dinnertime at either 5pm or 9pm). To measure individual differences in future-oriented
117 decision-making we used a standard monetary delay-discounting task. Our primary
118 hypothesis was that information about future meal timings would influence portion selection
119 at lunchtime. Specifically, we predicted that portion sizes would differ in each of the three
120 conditions and that participants would select smaller portions with a certain short IMI,
121 compared to a certain long IMI. Second, we proposed that steep money discounting would be
122 associated with IMI insensitivity in both certain and uncertain conditions. When the IMI was
123 certain, we hypothesized that steep discounters would show a smaller difference between
124 portions chosen at 5pm and 9pm. When the IMI was uncertain, we expected steep discounters
125 to select smaller portion sizes than the average of those chosen when meal times were certain.
126 Finally, to explore how BMI relates to future-oriented decision-making, we assessed
127 relationships between BMI, portion size, and monetary delay discounting.

128

129 **Method**

130 *Participants:* Participants ($N= 90$; 61 females, 29 males) had a mean age of 21.2y ($SD = 4.7$)
131 and were healthy staff or undergraduate and postgraduate students at the University of
132 Bristol, recruited through our laboratory volunteer database or as part of a course
133 requirement. They received either £5 (Sterling) or course credits in remuneration for their
134 assistance. The protocol was approved by the local Faculty of Science Human Research
135 Ethics Committee. *A priori*, we thought it was crucial that participants were familiar with the
136 foods we were including in the experiment. Therefore, we excluded fifteen participants who
137 indicated eating either of the test foods either 'never', or 'less than once a year'. A further
138 five participants were excluded for selecting the minimum portion of chow mein (20 kcal) for
139 lunch, in every condition. We suspect this reflects a technical error or otherwise a problem in
140 understanding the requirements of the tasks. Six participants had missing data for the delay-
141 discounting task due to a technical error. In these cases, values were entered as missing data.
142 The final dataset comprised 70 participants (46 females, 24 males), with a mean age of 21.0
143 years ($SD = 4.2$), and a mean BMI = 21.68 kg/m² ($SD = 2.6$; range = 16.6 - 27.1). In total, 7
144 participants were underweight, 55 participants were lean and 8 were overweight.

145

146 *Food images:* Based on previous research (Brunstrom, Collingwood, & Rogers, 2010) we
147 selected two different dishes that are commonly consumed as main meals in the UK: chicken
148 chow mein and chicken tikka masala with rice. For each dish, we photographed a series of
149 50 images with portion sizes ranging from 20 kcal to 1000 kcal, in equal 20-kcal steps. The
150 images were taken using a high-resolution digital camera under identical lighting conditions.
151 The meals were photographed on the same white plate (255-mm diameter).

152

153 *Measures*

154 *Liking:* Participants were shown a 400-kcal portion of the two test foods in a random order.

155 In each trial they responded on a 7-point scale with end anchor points labelled ‘extremely
156 dislike’ and ‘extremely like.’

157 *Familiarity:* Familiarity was assessed using a food-frequency questionnaire. Again,
158 participants were shown a 400-kcal portion of each food. In turn, they responded to the
159 question ‘How often do you eat this meal?’ by selecting one of the following options; ‘never,’
160 ‘less than once a year,’ ‘yearly,’ ‘every 2-3 months,’ ‘monthly,’ ‘weekly,’ or ‘daily.’ These
161 were coded 1-7 (least to most familiar).

162 *Appetite:* Measures of hunger and fullness were obtained using a 100-mm visual-analogue
163 rating scale headed ‘How [hungry/full/thirsty] do you feel right now?’, with end anchor
164 points ‘not at all’ and ‘extremely.’ All ratings were elicited on a computer.

165 *TFEQ:* Dietary behaviour was assessed using a computerised version of the 51-item Three
166 Factor Eating Questionnaire (TFEQ; (Stunkard & Messick, 1985). The instrument contains
167 36 items with a yes/no response format, 14 items on a 1-4 response scale and one vertical
168 rating. The relevant items were scored and aggregated into two scales. We were interested in
169 the Restraint and Disinhibition subscales. ‘Cognitive restraint’ (conscious control of food
170 intake to control body weight) and ‘disinhibition’ (loss of control over intake). Respectively,
171 higher scores indicate greater cognitive restraint and disinhibition. Internal-consistency
172 reliability coefficients (Cronbach’s α) were found to be above 0.70 and below 0.90 (de
173 Lauzon et al., 2004). The internal-consistency coefficient of the restraint and disinhibition
174 scales in the current study was 0.89.

175 *BMI:* To assess Body Mass Index (BMI), we measured participant’s height and weight at the
176 end of the experiment. BMI was calculated from measured weight/height².

177

178 *IMI portion task:* Two food images were presented on a VDU. We chose to use photographic
179 images as similar computer-based tasks have been shown to predict real food selection
180 (Pouyet, Cuvelier, Benattar, & Giboreau, 2015; Taylor, Yon, & Johnson, 2014). A fixed
181 portion (400 kcal) of chicken tikka masala was presented on the right and labelled ‘This meal
182 for dinner.’ A portion of chow mein was presented on the left and labelled ‘This meal for
183 lunch.’ The chow mein lunch portion could be increased or decreased by depressing the right
184 or left arrow-keys, respectively. In each trial the participants responded to the question ‘How
185 much would you eat for lunch RIGHT NOW if you had to eat all of the food on the right for
186 dinner at...[time inserted].’ In two of the trials the IMI was ‘certain.’ In one certain trial they
187 were told to expect their evening meal at 5pm. In the other they were told to expect it at 9pm.
188 In a third trial the IMI was ‘uncertain’ - they were told to expect the meal at either 5pm or
189 9pm. Participants completed a total of three trials. The order of the trials was randomised
190 across participants and each trial started with a randomly selected portion of chow mein.

191 To assess whether participants were more responsive to the uncertain future meal
192 times, we compared portions selected in the certain and uncertain conditions. The uncertain
193 IMI is framed around the same time points as the two certain IMIs (5pm and 9pm).
194 Therefore, the effect of uncertainty can be established by comparing portions chosen in the
195 uncertain condition with average of the portions chosen in the two certain condition.
196 Specifically, we used the three selected portion sizes (2 certain trials and 1 uncertain trial)
197 and computed a value (IMI index score) based on the following calculation: uncertain 5pm or
198 9pm - (certain 5pm + certain 9pm)/2). This provides a measure of the effect of uncertainty
199 (relative to certainty) on portion selection. We calculated a separate IMI index score for each
200 participant. A positive IMI index score indicates that larger portions were chosen in the
201 uncertain condition than in the average of the two portions selected in the certain conditions.

202

203 *Delay-discounting task:* Delay discounting was measured using a computerised forced-choice
204 task. The task was an adapted version of one previously introduced by Du and colleagues
205 (Du, Green, & Muerson, 2016). In a series of trials participants indicated whether they
206 preferred to receive a hypothetical delayed reward of £100 after a fixed interval (*e.g.*, 1 year)
207 or a smaller monetary amount immediately. Participants completed several blocks of 10
208 trials. In every trial the delayed reward was always £100. In the first trial of each block the
209 immediate reward was half the delayed value (£50). If the participant selected the immediate
210 reward, it was adjusted down to £16.66 (33.3% of its original value) in the second trial. If the
211 participant selected the delayed reward then it was adjusted up to £83.33 (the same difference
212 = £33.33). The same rationale was applied in subsequent trials (trials 3-10). However, in each
213 trial the adjustment amount decreased by 33.3% (*i.e.*, from £33.33 in trial 2 to £22.22 trial 3,
214 from £22.21 in trial 3 to £14.81 trial 4, and so on). This single ‘staircase’ approach
215 progressively converged around a point of indifference in which the delayed and immediate
216 amounts are equally likely to be selected.

217 Initially, three practice blocks were presented. In order, the hypothetical delays were 2
218 years, 1 year, and 6 months. This was followed by six further blocks. Each presented a
219 scenario with one of the following delays; 2 days, 7 days, 30 days, 6 months, 1 year, 2 years.
220 The order of these blocks was randomised across participants and responses were used to
221 calculate a measure of delay discounting. The delay-discounting task and the IMI portion task
222 were implemented using custom software (available on request) written in Visual Basic
223 (Microsoft version 6.0).

224 Following Myerson et al. (Myerson, Green, & Warusawitharana, 2001), for each
225 participant, a measure of delay discounting was obtained from area under the curve (AUC)
226 values derived from the delay-discounting task. AUC values were calculated using the
227 trapezoid method. Smaller AUC values indicate steeper discounting.

228

229 *Procedure:* Participants completed one 45-minute session between 12pm and 2pm. On arrival
230 they reported how long ago they last ate and then rated their appetite and thirst. They then
231 completed the IMI portion task, followed by liking and familiarity ratings, and then the delay-
232 discounting task. Finally, participants completed the TFEQ and we measured their BMI. At
233 the end of the study the participants were debriefed and thanked for their assistance.

234

235 *Data analysis:* First, to determine whether portion-size selection was influenced by
236 information about the IMI, we conducted a one-way, repeated-measures ANCOVA with
237 three conditions (portion size when the IMI was short, long and uncertain). We included
238 gender as a between-subjects factor and BMI and age as covariates. A paired *t*-test was used
239 to evaluate specific differences across participants between portion sizes chosen in the long
240 and short certain conditions. Second, to measure sensitivity to change in length of the certain
241 IMI, we assessed the difference between portions chosen in the two certain conditions. This
242 allowed us to calculate a Pearson's correlation to explore how certain IMI sensitivity related
243 to monetary delay discounting. Similarly, we calculated the correlation between delay
244 discounting and sensitivity to the uncertain IMI, relative to the certain IMIs (IMI index
245 score). In addition, we assessed correlations between BMI and both IMI index score and
246 delay discounting.

247

248 Post-hoc analyses were conducted to investigate whether individual differences in
249 delay discounting mediated the relationship between BMI and portion-size selection in
250 uncertain IMIs. For a mediating relationship to be confirmed, four key criteria must be met.
251 Criterion 1, the independent variable (IV) and the dependent variable (DV) must be
252 significantly associated (Baron & Kenny, 1986). Criterion 2, the IV and the mediator must be
significantly associated; Criterion 3, the mediator and the DV must be significantly

253 associated; Criterion 4, when the mediator is controlled for in a regression of the IV on the
 254 DV, the β -value relating the IV to the DV becomes insignificant. In our post-hoc analysis, we
 255 entered the IMI index scores as the IV, BMI as the DV, and impulsivity as the mediator.

256 All four criterion were explored using multiple regression analysis. The
 257 unstandardized regression coefficients and standard errors of the relationship between the IV
 258 and the mediator, and between the DV and the mediator, are used to calculate the path
 259 coefficient (b_{ab}) and its standard error ($^s b_{ab}$). The path coefficient is divided by the standard
 260 error to give a t -ratio. If the t -ratio exceeds ± 1.96 , then the indirect path is significant and a
 261 mediating relationship is confirmed. All data were analysed using IBM SPSS statistics
 262 version 21 (IBM, New York, USA).

263 Results

264 *Participant characteristics:* Table 1 shows mean scores for liking, appetite, TFEQ, and
 265 familiarity, as well as participant characteristics. Both BMI and Delay discounting AUC
 266 scores were not related to liking, hunger, fullness, familiarity restraint or disinhibition (See
 267 Table 2). Mean TFEQ-restraint score ($M = 6.7$, $SD = 3.6$) and mean TFEQ-disinhibition score
 268 ($M = 6.3$, $SD = 2.6$) were all in the low range (Lesdema et al., 2012; Stunkard & Messick,
 269 1985).

270

271 Table 1. Means and standard deviations (SD) for participant characteristics, questionnaires,
 272 ratings and delay discounting AUC

Measure (units/range)	Mean (SD)	Range (min-max)
Age (y)	21.0 (4.2)	18.0 – 43.0
BMI (kg/m ²)	21.7 (2.6)	16.7 – 27.1

TFEQ-restraint (0 - 21)	6.7 (3.6)	1.0 – 17.0
TFEQ-disinhibition (0 - 16)	6.3 (2.6)	1.0 – 13.0
Delay discounting (AUC)	0.6 (0.2)	0.0 – 1.0
Appetite (1-7)	5.0 (1.73)	1.0 – 7.0
Familiarity (chicken tikka and chow mein; (2-14)	9.8 (1.33)	2.0 – 14.0

273 (N = 70; 46 female, 24 male)

274

275 Table 2. Relationships (Pearson's correlations) between inter-meal interval (IMI) index score,
276 delay discounting area under the curve (DD AUC), TFEQ, BMI, liking, hunger, and fullness.

	1	2	3	4	5	6	7	8
1. IMI index								
2. DD AUC	.29*							
3. TFEQ-Disinhibition	.18	.17						
4. TFEQ-Restraint	-.01	-.03	.13					
5. BMI	-.27*	-.40**	-.16	.29*				
6. Liking	-.20	-.13	-.09	-.11	.04			
7. Fullness	.16	.11	.02	.03	.14	-.07		
8. Hunger	-.139	.051	-.03	-.12	-.08	.11	-.74**	

277 * $p < .05$

278 ** $p < .01$

279

280 *IMI portion task:* Our analysis revealed a main effect of IMI on portion selection after
281 controlling for age, gender and BMI, $F(2,132) = 4.53, p = .012, \eta^2 = .06$. Specifically,
282 participants chose larger portions with a certain long IMI (dinner at 9pm; $M = 549.1$ kcal, $SD,$
283 205.3) than a short certain IMI (dinner at 5pm; $M = 423.4$ kcal, $SD = 217.1$), $t(69) = 6.02, p$
284 $= .00$. Covariates, age, gender and BMI did not predict variance in portion selection (all $p >$
285 $.05$). Correlations between IMI index score and liking, fullness, TFEQ-restraint and TFEQ-
286 disinhibition failed to reach significance (see Table 2).

287 *Relationship between discounting and IMI sensitivity:* There was a medium sized, but non-
288 statistically significant, correlation between delay discounting AUC and the difference
289 between portion size at long and short certain IMIs, $r(62) = .18, p = .15^1$. Consistent with our
290 hypothesis, we found a significant positive correlation between delay discounting AUC and
291 IMI index score, $r(62) = .29, p < .05^1$. Participants who exhibited steeper discounting (lower
292 AUC) chose smaller portions when the IMI was uncertain than when it was certain (See
293 Supplemental Material for visual representation of relationship between IMI index and delay
294 discounting).

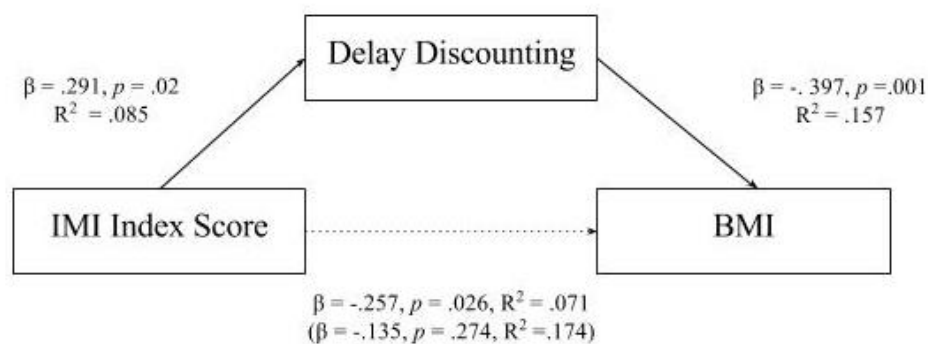
295 *Relationship between BMI with discounting and portion size selection at the uncertain IMI:*
296 There was a significant negative correlation between BMI and IMI index score $r(69) = -$
297 $.27, p < .05$. Individuals with a high BMI chose smaller portions when the IMI was uncertain,
298 compared to when it was certain. There was also a significant negative correlation between
299 BMI and delay discounting AUC $r(62) = -.40, p < .001$. Participants who showed steeper
300 discounting had a higher BMI than those with shallower discounting. Relationship between

¹ Degrees of freedom differ due to missing data

301 BMI and inter-meal-interval (IMI) index score. (See Supplemental Material for visual
 302 representation of relationship between IMI index and BMI).

303

304 *Post-hoc mediation analysis:* Significant relationships were confirmed between IMI index
 305 score and BMI (criterion 1), between delay discounting AUC and IMI index score (criterion
 306 2) and between delay discounting AUC and BMI (criterion 3). When delay discounting AUC
 307 was controlled for in a regression of IMI index score on BMI, IMI index score no longer
 308 predicted BMI (criterion 4). Figure 1 shows the regression coefficients associated with tests
 309 of each relationship. Subsequently, the Sobel test (Sobel, 1982) confirmed that the two-tailed
 310 mediator was significant, $t(62) = -2.59, p = .012$. As all criteria for mediation were met and
 311 the Sobel test was significant, this suggests that delay discounting mediates the relationship
 312 between BMI and smaller portion size selection at the uncertain, relative to certain, IMI



313

314 *Figure 1.* Delay discounting as a mediator of the relationship between selection of smaller
 315 portion sizes at the uncertain inter-meal interval (IMI index score) and BMI. Unstandardized
 316 β , p and R^2 values are shown for each relationship. Regression coefficients associated with
 317 Criterion 4 (when the mediator is controlled for in a regression of IMI index score on delay
 318 discounting) are shown in brackets.

319

320 *Post-hoc power calculation:* To assess satisfactory statistical power, we conducted a post hoc
321 power analysis. The medium effect size states that we were underpowered to detect an
322 association between delay discounting and the difference between portion sizes selected at
323 the certain IMIs. The calculation revealed a sample size of 240 would be required to detect
324 this effect with an α of 0.05 and a $1-\beta$ of 0.80.

325

326

Discussion

327 This study assessed how information about IMIs influences portion size decisions and
328 whether steep delay discounters respond differently to the predictability of an IMI. Our
329 primary hypothesis was that information about future IMIs would influence portion size
330 decisions. Secondly, we hypothesised that steep monetary delay discounters would be less
331 sensitive to information about the duration of the certain IMIs, and show a small differences
332 between portions selected in the long and short IMIs. In particular, we predicted steep
333 discounters would show even greater disregard for future meal times in the uncertain IMI.

334 Consistent with our first hypothesis, participants chose larger portions in response to
335 the certain long IMI than in response to the certain short IMI. This is the first demonstration
336 that people use information about future meal timings to make in-the-moment decisions about
337 how much to eat. Greater monetary delay discounting was associated with smaller portion
338 selection in response to the uncertain IMI, compared to the average of those chosen in the
339 certain IMIs. We suggest that shallow discounters selected larger portions to protect against
340 possible hunger during the IMI. Consistent with our hypothesis, steep delay discounters
341 appeared to show a disregard for the uncertain IMI, possibly due to a lack of concern for
342 potential hunger between meals. However, steep and shallow discounters selected similar
343 portion sizes when the IMI was certain, suggesting that delay discounting is less relevant

344 when an IMI is known. Consistent with this idea, individuals show greater discounting of a
345 future reward when the occurrence of a delayed event is less certain (Baumann & Odum,
346 2012; Green & Myerson, 2010; Patak & Reynolds, 2007). Our results suggest that variability
347 in the timing of the event also increases discounting. In the future, researchers should
348 differentiate between irregular eating in the presence or absence of uncertainty. These
349 observations suggest that dietary discounting is more likely to be expressed when meal times
350 are uncertain. Hence, a distinction between certain and uncertain meal timings might be
351 helpful, especially in studies seeking to understand relationships between chaotic eating,
352 discounting, and BMI.

353 We also predicted that steep discounters would be less likely to plan their meals based
354 on the duration of the certain IMI. In line with this, delay discounting was associated with a
355 smaller difference between portions selected at the long and short certain IMIs. This suggests
356 that steep discounters were less sensitive to information about future meal timings, whereas
357 future-oriented individuals were more likely to plan for the IMI. Although this relationship
358 failed to reach statistical significance, the effect sizes indicate a small-to-medium sized
359 association, suggesting that the current study was potentially underpowered (a sample size of
360 240 would be required to detect this effect, with an α of 0.05 and a $1-\beta$ of 0.80).

361 Temporal discounting is generally regarded as a trait that promotes overconsumption.
362 Our data show that delay discounting might actually reduce self-selected portion size.
363 Specifically, the expression and downstream effects of discounting might depend upon
364 whether a meal is planned and whether an IMI is certain or uncertain. These findings could
365 help to explain previous inconsistent associations between delay discounting and eating
366 behaviour. Dietary discounting is typically conceptualised as a trade-off between immediate
367 food reward and long-term future health costs. Our data suggests that discounting is also
368 expressed in shorter-term delays from one meal to the next. These distinctions are subtle yet

369 potentially essential, and are generally overlooked in studies exploring the acute effects of
370 temporal discounting on food intake. A more nuanced understanding of how meal timings
371 influence future-oriented decisions will contribute to the development of an evidence base,
372 which can inform guidelines around structured eating and meal planning.

373 Our post-hoc analysis suggests that delay discounting mediated a relationship between
374 having a higher BMI and selecting *smaller* portions with an uncertain IMI. This appears
375 counterintuitive; steep discounters had higher BMIs, yet chose smaller portions. One
376 possibility is that a lack of concern for future hunger promotes various compensatory
377 behaviours, such as the selection of energy-dense snacks between meals. In line with this,
378 both chaotic eating and impulsivity have been associated with a greater tendency to snack
379 between meals (Fay, et al., 2015) and also greater consumption of palatable foods (Lumley,
380 2016). Further research is required to determine whether snacking behaviour is more
381 prevalent in individuals who less sensitive to information about IMIs.

382 The study may be limited by using computer-based judgements of food decisions.
383 Nevertheless, our focus was to understand relationships between discounting and meal
384 planning. Although computer-based portion judgments are shown to be predictive of real
385 food intake (Pouyet, et al., 2015; Taylor, et al., 2014), it remains to be determined whether
386 the same relationships might be observed in a study of food intake. This was beyond the
387 scope of the present study but might be considered in future research. Additionally, as
388 participants were university students with a relatively narrow range of BMIs, the
389 generalizability of our findings remains unclear. The generalisability of our conclusions that
390 delay discounters are less sensitive to information about future meal timings are somewhat
391 limited by the lack of statistical power limited our conclusions; subsequent research is
392 required to explore these relationships in a larger and more representative sample. Finally, as

393 mood is shown to influence delay discounting (Koff & Lucas, 2011), subsequent studies
394 could assess how mood influences decision-making regarding discounting of meal timings.

395

396 *Concluding remarks*

397 In summary, steep delay discounters selected smaller portions in response to an
398 uncertain IMI, compared to the certain IMIs. We reasoned that in conditions of uncertainty,
399 non-future oriented individuals were less concerned with potential hunger or fullness between
400 meals and selected how much they would like in the moment. These results suggest that delay
401 discounting is more likely to be expressed in a ‘chaotic’ eating environment. Future studies
402 are required to assess these relationships in a wider sample and with real food intake to
403 improve generalizability of our conclusions. Our findings merit consideration because they
404 demonstrate how short-term discounting can influence portion-size decisions.

405

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412

413 **Conflict of interest.**

414 The authors declare no conflict of interest.

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References

- 419 Appelhans, B. M., Woolf, K., Pagoto, S. L., Schneider, K. L., Whited, M. C., & Liebman, R.
420 (2011). Inhibiting food reward: delay discounting, food reward sensitivity, and
421 palatable food intake in overweight and obese women. *Obesity*, *19*(11), 2175-2182.
422 doi: 10.1038/oby.2011.57
- 423 Barlow, P., Reeves, A., McKee, M., Galea, G., & Stuckler, D. (2016). Unhealthy diets,
424 obesity and time discounting: a systematic literature review and network analysis.
425 *Obesity Reviews*, *17*(9), 810-819. doi: 10.1111/obr.12431
- 426 Baron, R. M., & Kenny, D. A. (1986). The moderator–mediator variable distinction in social
427 psychological research: Conceptual, strategic, and statistical considerations. *Journal*
428 *of Personality and Social Psychology*, *51*(6), 1173. doi: 10.1037/0022-3514.51.6.1173
- 429 Baumann, A. A., & Odum, A. L. (2012). Impulsivity, Risk Taking, and Timing. *Behavioural*
430 *Processes*, *90*(3), 408-414. doi: 10.1016/j.beproc.2012.04.005
- 431 Berg, C., & Forslund, H. B. (2015). The Influence of Portion Size and Timing of Meals on
432 Weight Balance and Obesity. *Current Obesity Reports*, *4*(1), 11-18. doi:
433 10.1007/s13679-015-0138-y
- 434 Brunstrom, J. M., Collingwood, J., & Rogers, P. J. (2010). Perceived volume, expected
435 satiation, and the energy content of self-selected meals. *Appetite*, *55*(1), 25-29. doi:
436 <http://dx.doi.org/10.1016/j.appet.2010.03.005>
- 437 Brunstrom, J. M., & Rogers, P. J. (2009). How many calories are on our plate? Expected
438 fullness, not liking, determines meal-size selection. *Obesity (Silver Spring)*, *17*(10),
439 1884-1890. doi: 10.1038/oby.2009.201
- 440 Brunstrom, J. M., Shakeshaft, N. G., & Scott-Samuel, N. E. (2008). Measuring ‘expected
441 satiety’ in a range of common foods using a method of constant stimuli. *Appetite*,
442 *51*(3), 604-614. doi: <http://dx.doi.org/10.1016/j.appet.2008.04.017>
- 443 da Matta, A., Gonçalves, F. L., & Bizarro, L. (2012). Delay discounting: Concepts and
444 measures. *Psychology & Neuroscience*, *5*(2), 135-146. doi: 10.3922/j.psns.2012.2.03
- 445 de Lauzon, B., Romon, M., Deschamps, V., Lafay, L., Borys, J. M., Karlsson, J., . . . Charles,
446 M. A. (2004). The Three-Factor Eating Questionnaire-R18 is able to distinguish
447 among different eating patterns in a general population. *The Journal of Nutrition*,
448 *134*(9), 2372-2380.
- 449 Du, W., Green, L., & Muerson, J. (2016). Cross-cultural comparisons of discounting delayed
450 and probabilistic rewards. *The Psychological Record*, *52*(4), 479-493.
- 451 Eisenstein, S. A., Gredysa, D. M., Antenor-Dorsey, J. A., Green, L., Arbelaez, A. M., Koller,
452 J. M., . . . Hershey, T. (2015). Insulin, Central Dopamine D2 Receptors, and Monetary
453 Reward Discounting in Obesity. *PLoS One*, *10*(7), e0133621. doi:
454 10.1371/journal.pone.0133621
- 455 Evenden, J. L. (1999). Varieties of impulsivity. *Psychopharmacology (Berl)*, *146*(4), 348-
456 361.
- 457 Fay, S. H., Ferriday, D., Hinton, E. C., Shakeshaft, N. G., Rogers, P. J., & Brunstrom, J. M.
458 (2011). What determines real-world meal size? Evidence for pre-meal planning.
459 *Appetite*, *56*(2), 284-289. doi: <http://dx.doi.org/10.1016/j.appet.2011.01.006>
- 460 Fay, S. H., White, M. J., Finlayson, G., & King, N. A. (2015). Psychological predictors of
461 opportunistic snacking in the absence of hunger. *Eating Behaviors*, *18*, 156-159. doi:
462 10.1016/j.eatbeh.2015.05.014

- 463 Green, L., & Myerson, J. (2010). Experimental and correlational analyses of delay and
464 probability discounting. In: Madden GJ, Bickel WK (eds) *Impulsivity: the behavioral*
465 *and neurological science of discounting. American Psychological Association*, 67-92.
466 doi: 10.1037/12069-003
- 467 Gregorios-Pippas, L., Tobler, P. N., & Schultz, W. (2009). Short-term temporal discounting
468 of reward value in human ventral striatum. *Journal of Neurophysiology*, 101(3), 1507-
469 1523. doi: 10.1152/jn.90730.2008
- 470 Greville, W. J., & Buehner, M. J. (2010). Temporal predictability facilitates causal learning.
471 *Journal of Experimental Psychology: General*, 139(4), 756. doi: 10.1037/a0020976
- 472 Hendrickson, K. L., & Rasmussen, E. B. (2013). Effects of mindful eating training on delay
473 and probability discounting for food and money in obese and healthy-weight
474 individuals. *Behaviour Research and Therapy*, 51(7), 399-409. doi:
475 10.1016/j.brat.2013.04.002
- 476 Hendrickson, K. L., Rasmussen, E. B., & Lawyer, S. R. (2015). Measurement and validation
477 of measures for impulsive food choice across obese and healthy-weight individuals.
478 *Appetite*, 90, 254-263. doi: 10.1016/j.appet.2015.03.015
- 479 Koff, E., & Lucas, M. (2011). Mood moderates the relationship between impulsiveness and
480 delay discounting. *Personality and Individual Differences*, 50(7). doi:
481 10.1016/j.paid.2011.01.016
- 482 Kulendran, M., Vlaev, I., Sugden, C., King, D., Ashrafian, H., Gately, P., & Darzi, A. (2014).
483 Neuropsychological assessment as a predictor of weight loss in obese adolescents.
484 *International Journal of Obesity*, 38(4), 507-512. doi: 10.1038/ijo.2013.198
- 485 Leitch, M. A., Morgan, M. J., & Yeomans, M. R. (2013). Different subtypes of impulsivity
486 differentiate uncontrolled eating and dietary restraint. *Appetite*, 69, 54-63. doi:
487 10.1016/j.appet.2013.05.007
- 488 Lesdema, A., Fromentin, G., Daudin, J. J., Arlotti, A., Vinoy, S., Tome, D., & Marsset-
489 Baglieri, A. (2012). Characterization of the Three-Factor Eating Questionnaire scores
490 of a young French cohort. *Appetite*, 59(2), 385-390. doi: 10.1016/j.appet.2012.05.027
- 491 Lumley, J., Stevenson, RJ, Oaten, MJ, Mahmut, M, Yeomans, MR. (2016). Individual
492 differences in impulsivity and their relationship to a Western-style diet. *Personality*
493 *and Individual Differences*, 97, 178-185. doi: 10.1016/j.paid.2016.03.055
- 494 MacKillop, J., Amlung, M. T., Few, L. R., Ray, L. A., Sweet, L. H., & Munafò, M. R.
495 (2011). Delayed reward discounting and addictive behavior: a meta-analysis.
496 *Psychopharmacology (Berl)*, 216(3), 305-321. doi: 10.1007/s00213-011-2229-0
- 497 Manwaring, J. L., Green, L., Myerson, J., Strube, M. J., & Wilfley, D. E. (2011). Discounting
498 of various types of rewards by women with and without binge eating disorder:
499 Evidence for general rather than specific differences. *The Psychological Record*,
500 61(4), 561-582.
- 501 Mazur, J. E. (2001). Hyperbolic value addition and general models of animal choice.
502 *Psychological Review*, 108(1), 96-112.
- 503 McClure, S. M., Ericson, K. M., Laibson, D. I., Loewenstein, G., & Cohen, J. D. (2008). Time
504 Discounting for Primary Rewards. *The Journal of Neuroscience*, 27(21), 5796-5804.
505 doi: 10.1523/jneurosci.4246-06.2007
- 506 Moeller, F. G., Barratt, E. S., Dougherty, D. M., Schmitz, J. M., & Swann, A. C. (2001).
507 Psychiatric aspects of impulsivity. *American Journal of Psychiatry*, 158(11), 1783-
508 1793. doi: 10.1176/appi.ajp.158.11.1783
- 509 Murata, M. (2000). Secular trends in growth and changes in eating patterns of Japanese
510 children. *American Journal of Clinical Nutrition*, 72(5), 1379-1383.

- 511 Myerson, J., Green, L., & Warusawitharana, M. (2001). Area under the curve as a measure of
512 discounting. *Journal of the Experimental Analysis of Behavior*, *76*(2), 235-243. doi:
513 10.1901/jeab.2001.76-235
- 514 Nederkoorn, C., Houben, K., Hofmann, W., Roefs, A., & Jansen, A. (2010). Control yourself
515 or just eat what you like? Weight gain over a year is predicted by an interactive effect
516 of response inhibition and implicit preference for snack foods. *Health Psychology*,
517 *29*(4), 389-393. doi: 10.1037/a0019921
- 518 Patak, M., & Reynolds, B. (2007). Question-based assessments of delay discounting: do
519 respondents spontaneously incorporate uncertainty into their valuations for delayed
520 rewards? *Addictive behaviors*, *32*(2), 351-357. doi: 10.1016/j.addbeh.2006.03.034
- 521 Pouyet, V., Cuvelier, G., Benattar, L., & Giboreau, A. (2015). A photographic method to
522 measure food item intake. Validation in geriatric institutions. *Appetite*, *84*, 11-19. doi:
523 10.1016/j.appet.2014.09.012
- 524 Rasmussen, E. B., Lawyer, S. R., & Reilly, W. (2010). Percent body fat is related to delay
525 and probability discounting for food in humans. *Behavioural Processes*, *83*(1), 23-30.
526 doi: 10.1016/j.beproc.2009.09.001
- 527 Rollins, B. Y., Dearing, K. K., & Epstein, L. H. (2010). Delay discounting moderates the
528 effect of food reinforcement on energy intake among non-obese women. *Appetite*,
529 *55*(3), 420-425. doi: 10.1016/j.appet.2010.07.014
- 530 Samuelson, G. (2000). Dietary habits and nutritional status in adolescents over Europe. An
531 overview of current studies in the Nordic countries. *European Journal of Clinical*
532 *Nutrition*, *54*(3), S21-28. doi: <http://dx.doi.org/10.1038/sj.ejcn.1600980>
- 533 Shelley, M. K. (1993). Outcome Signs, Question Frames and Discount Rates. *Management*
534 *Science*. Jul1993, *39*(7), 806. doi: <http://dx.doi.org/10.1287/mnsc.39.7.806>
- 535 Sierra-Johnson, J., Unden, A. L., Linstrand, M., Rosell, M., Sjogren, P., Kolak, M., . . .
536 Hellenius, M. L. (2008). Eating meals irregularly: a novel environmental risk factor
537 for the metabolic syndrome. *Obesity (Silver Spring)*, *16*(6), 1302-1307. doi:
538 10.1038/oby.2008.203
- 539 Sobel, M. (1982). Asymptotic confidence intervals for indirect effects in structural equation
540 models. *Sociological Methodology*, *13*, 290-312.
- 541 Stoeckel, L. E., Murdaugh, D.L., Cox, J.E., Cook, E.W., Weller, R.E. (2013). Greater
542 impulsivity is associated with decreased brain activation in obese women during a
543 delay discounting task. *Brain Imaging and Behavior*, *7*(2), 116-128.
- 544 Stojek, M. M., Fischer, S., Murphy, C. M., & MacKillop, J. (2014). The role of impulsivity
545 traits and delayed reward discounting in dysregulated eating and drinking among
546 heavy drinkers. *Appetite*, *80*, 81-88. doi: 10.1016/j.appet.2014.05.004
- 547 Stunkard, A. J., & Messick, S. (1985). The three-factor eating questionnaire to measure
548 dietary restraint, disinhibition and hunger. *Journal of Psychosomatic Research*, *29*(1),
549 71-83. doi: [http://dx.doi.org/10.1016/0022-3999\(85\)90010-8](http://dx.doi.org/10.1016/0022-3999(85)90010-8)
- 550 Tanaka, S. C., Doya, K., Okada, G., Ueda, K., Okamoto, Y., & Yamawaki, S. (2004).
551 Prediction of immediate and future rewards differentially recruits cortico-basal
552 ganglia loops. *Nature Neuroscience*, *7*(8), 887-893. doi: doi:10.1038/nn1279
- 553 Taylor, J. C., Yon, B. A., & Johnson, R. K. (2014). Reliability and validity of digital imaging
554 as a measure of schoolchildren's fruit and vegetable consumption. *Journal of the*
555 *Academy of Nutrition and Dietetics*, *114*(9), 1359-1366. doi:
556 10.1016/j.jand.2014.02.029
- 557 Verplanken, B., & Sato, A. (2011). The Psychology of Impulse Buying: An Integrative Self-
558 Regulation Approach. [article]. *Journal of Consumer Policy*, *34*(2), 197-120. doi:
559 <http://dx.doi.org/10.1007%2Fs10603-011-9158-5>

- 560 Warde, A., & Yates, L. (2016). Understanding Eating Events: Snacks and Meal Patterns in
561 Great Britain. *Food, Culture and Society*, 1-22. doi: 1243763
- 562 Weller, R. E., Cook, E. W., Avsar, K. B., & Cox, J. E. (2008). Obese women show greater
563 delay discounting than healthy-weight women. *Appetite*, 51(3), 563-569. doi:
564 <http://dx.doi.org/10.1016/j.appet.2008.04.010>
- 565 Whiteside, S. P., & Lynam, D. R. (2016). The Five Factor Model and impulsivity: using a
566 structural model of personality to understand impulsivity. *Personality and Individual*
567 *Differences*, 30(4), 669-689. doi: 10.1016/S0191-8869(00)00064-7
- 568 Wilkinson, L. L., Hinton, E. C., Fay, S. H., Ferriday, D., Rogers, P. J., & Brunstrom, J. M.
569 (2012). Computer-based assessments of expected satiety predict behavioural measures
570 of portion-size selection and food intake. *Appetite*, 59(3), 933-938. doi:
571 10.1016/j.appet.2012.09.007
- 572 Zhang, L., & Rashad, I. (2008). Obesity and time preference: the health consequences of
573 discounting the future. *Journal of Biosocial Science*, 40(1), 97-113. doi:
574 10.1017/s0021932007002039
- 575